## Bayesian Statistics

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# GP-LVM for density estimation (Kundu & Dunson, 2013, Biometrika)

Consider the latent variable model,

$$y_i = \mu(\eta_i) + \epsilon_i, \quad \epsilon_i \sim N(0, \sigma^2), (i = 1, ..., n)$$
  
 $\mu \sim GP(0, c), \quad \sigma \sim IG(a, b), \quad \eta_i \sim U(0, 1),$ 

- $\eta_i$ 's are subject specific latent variables,  $\mu \in C[0, 1]$  is a *transfer function* relating the latent variables to the observed variables.
- The density of y conditional on the transfer function μ and scale σ is obtained on marginalizing out the latent variable as

$$f(y;\mu,\sigma) \stackrel{\text{def}}{=} f_{\mu,\sigma}(y) = \int_0^1 \phi_\sigma \{y-\mu(x)\} dx \approx \frac{1}{T} \sum_{j=1}^T \phi_\sigma \Big\{y-\mu\left(\frac{j}{T}\right)\Big\}.$$

### Flexibility of GP-LVM

• Let 
$$F_0(t) = \int_{-\infty}^t f_0(x) dx$$
.

• Letting  $\mu_0(x) = F_0^{-1}(x)$ , one obtains

$$f_{\mu_0,\sigma}(y) = \int_0^1 \phi_\sigma(y - F_0^{-1}(x)) dx = \int_{-\infty}^\infty \phi_\sigma(y - t) f_0(t) dt,$$

where the second equality follows from the change of variable theorem.

- $\blacktriangleright \|\phi_{\sigma} * f_0 f_0\| \to 0$
- Posterior computation: i) μ | y, η, σ is Gaussian ii) η | y, μ, σ
   Metropolis Hastings iii) σ | y, η, σ is Inverse Gamma

# Some applications of the Dirichlet process Mixture models

- $\theta_i$  = random effects specific to "subject" *i*
- Hierarchical models let  $\theta_i \sim P$
- P = random effects distribution
- Choice of P critical in controlling borrowing of information

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- Meta Analysis: combine data from multiple studies to make overall conclusion (e.g., drug is effective)
- Multi-level Designs: subjects are nested in schools, regions or study centers

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 Longitudinal Data: data collected for subject over time important to accommodate within-subject dependence

### Joint modeling of data from different domains

- 1. Images and captions
- 2. Diagnostic images or functional predictors & health responses

- 3. Multiple types of omics data (sequence & expression)
- Multi-task learning: borrow strength across tasks
  - 1. Multiple images, music pieces, security videos
  - 2. Compressive sensing
  - 3. User preferences in different domains (film, books, etc)

- Increasing concern about adverse effects of environmental estrogens on human development
- Rodent uterotrophic bioassay: system for identifying suspected agonists or antagonists of estrogen.

- <u>OECD study</u>: collected data from 19 laboratories to investigate consistency of effects of known agonist (EE) & antagonist (ZM)
- y<sub>ij</sub> = uterus weight for rat j in lab i
- ► x<sub>ii</sub>=lab indicator, dose of EE, dose of ZM

Can potentially fit normal random effects model,

$$\begin{aligned} \mathbf{y}_{ij} &= \mathbf{x}'_{ij} \mathbf{b}_i + \epsilon_{ij}, \epsilon_{ij} \sim \mathsf{N}(\mathbf{0}, \sigma^2) \\ \boldsymbol{\theta}_i &\sim \mathsf{N}_{\mathsf{P}}(\boldsymbol{\theta}, \boldsymbol{\Sigma}) \end{aligned}$$

 Normal distribution has light tails & does not allow outlying labs or clusters of labs

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- Conclusions may be sensitive to violations of normality
- Appealing to have a more flexible approach available

A simple case corresponds to the linear mixed effects model

$$y_{ij} = \mathbf{x}_{ij}\boldsymbol{\beta} + \mathbf{z}_{ij}\mathbf{b}_i + \epsilon_{ij}, \epsilon_{ij} \sim \mathsf{N}(0, \sigma^2)$$
  
$$\mathbf{b}_i \sim P, \quad P \sim DP(\alpha P_0)$$

- ▶ DP prior on *P*, the distribution of the random effects
- Useful semiparametric model for longitudinal & correlated data
- Bush & MacEachern (1996), Müller & Rosner (1997), Kleinman & Ibrahim (1998), Ishwaran & Takahara (2002), etc

## Application 2 - longitudinal data (log PdG trajectories (Bigelow & Dunson, JASA, 08))



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- Interest in estimating a collection of functions,  $\{f_i\}_{i=1}^n$ .
- Longitudinal trajectories for different individuals
- We will focus on the following model:

$$egin{array}{rcl} y_{ij}&=&f_i(t_{ij})+\epsilon_{ij}, & \epsilon_{ij}\sim t_
u(\sigma^2)\ f_i(t)&=&\sum_{j=1}^peta_{ij}b_j(t)=\mathbf{b}(t)'eta_i\ eta_j&\sim&P \end{array}$$

**b** = {*bj*}=basis functions,  $\beta_i$  =basis coefficients

 Subject-specific basis coefficients, β<sub>i</sub>, allow variability in the functional trajectories for different individuals

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- Heterogeneity among subjects controlled by the random effects distribution, *P*
- Number of basis functions, p, is not small ( $p \ge 20$ )

- Characterize variability in growth curves & cluster subjects having similar trajectories
- Can be accomplished using DPM linear mixed model with

$$f_i(t_{ij}) = \sum_{l=1}^{p} \beta_{il} b_l(t_{ij})$$
$$\beta_i \sim P = \sum_{h=1}^{\infty} \pi_h \delta_{\beta_h^*}$$

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### **Comments- Functional Dirichlet Process**

Recalling the DP stick-breaking property (Sethuraman, 1994):

$$eta_i \sim P = \sum_{h=1}^{\infty} V_h \prod_{l < h} (1 - V_l) \delta_{eta_h^*}, V_h \overset{i.i.d}{\sim} \mathsf{Beta}(1, \alpha), eta_h^* \sim P_0$$

- Hence, the *n* subjects are grouped into  $k \leq n$  clusters
- Subjects in cluster *I* all have  $\beta_i = \beta_i^*$ .
- Provides a semiparametric Bayes version of latent trajectory class or growth mixture models.

Avoids fixing the number of clusters in advance

### Comments- Functional Dirichlet Process

- The curve in cluster *I* is  $f(t) = \mathbf{b}(t)'\beta_I^*$ .
- The number of functional clusters in n growth curves is treated as unknown
- Gibbs samplers are straightforward to generalize
- Number of clusters and configuration of subjects into clusters varies across the MCMC iterations

Problem: label switching!

- Problem arises because the labels on the cluster-specific parameters are ambiguous, so vary in meaning across the iterations
- Not meaningful to calculate posterior summaries of ??h across the iterations
- Strategies:
  - 1. Relabeling algorithms that align the clusters after running MCMC (Stephens, 00);
  - 2. Define clusters as individuals that are grouped together with high posterior probability

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- 3. Estimate optimal clustering (Dahl, 06; Lau & Green, 97)
- 4. Ignore problem & avoid cluster-specific inferences