- 6.5 sem =  $\frac{0.5}{\sqrt{40}}$  = 0.079 for normal men and  $\frac{0.4}{\sqrt{32}}$  = 0.071 for men with chronic airflow limitation.
- It means that the distribution of mean triceps skin-fold thickness from repeated samples of size 40 drawn from the population of normal men can be considered to be normal with mean  $\mu$  and variance  $\frac{\sigma^2}{n} = \frac{s^2}{n} = \frac{0.5^2}{40} = 0.0063$ . A similar statement holds for men with chronic airflow limitation.
- **6.7** 2.583
- **6.8** -1.313
- **6.9** 2.365
- We refer to Table 6. The lower 2.5th percentile is 0.0506 and is denoted by  $\chi^2_{2,025}$ . The upper 2.5th percentile is 7.38 and is denoted by  $\chi^2_{2,975}$ .
- 6.11 We have that  $\bar{x} = \frac{215}{25} = 8.6$  days. Therefore, a 95% confidence interval for  $\mu$  is given by

$$\overline{x} \pm \frac{t_{24, .975}s}{\sqrt{n}} = 8.6 \pm \frac{2.064(5.72)}{\sqrt{25}}$$
  
= 8.6 ± 2.36 = (6.24, 10.96)

The 95% confidence interval is computed from  $\overline{x} \pm t_{n-1,.975} \frac{s}{\sqrt{n}}$ . We have that  $\overline{x} = 7.84$ , s = 3.21. There-fore, we have the following 95% confidence interval

$$7.84 \pm t_{24,975} \times \frac{3.21}{\sqrt{25}} = 7.84 \pm 2.064 \times \frac{3.21}{5}$$
$$= 7.84 \pm 1.33 = (6.51, 9.17)$$

**6.13** A 90% confidence interval is given by

$$\overline{x} \pm t_{n-1,95} \frac{s}{\sqrt{n}} = 7.84 \pm t_{24,95} \times \frac{3.21}{\sqrt{25}}$$
$$= 7.84 \pm 1.711 \times \frac{3.21}{5}$$
$$= 7.84 \pm 1.10 = (6.74, 8.94)$$

- The 90% confidence interval should be shorter than the 95% confidence interval, since we are requiring less confidence. This is indeed the case.
- 6.15 Our best estimate is given by  $\hat{p} = \frac{11}{25} = .44$ .

- **6.16** The standard error =  $\sqrt{\frac{\hat{p}\hat{q}}{n}} = \sqrt{\frac{.44 \times .56}{25}} = .099$ .
- Since  $n\hat{p}\hat{q} = 25 \times .44 \times .56 = 6.16 \ge 5$ , we can use the normal theory method. Therefore, a 95% confidence interval for the percentage of males discharged from Pennsylvania hospitals is given by

$$\hat{p} \pm 1.96 \sqrt{\frac{\hat{p}\hat{q}}{n}} = .44 \pm 1.96(.099)$$
$$= .44 \pm .195$$
$$= (.25, .63)$$

- 6.40 The best point estimate of p is  $\hat{p} = \frac{20}{100} = .2$ .
- We can compute the 95% confidence interval for p based on the normal approximation to the binomial  $(n\hat{p}\hat{q} = 16 \ge 5)$  and note if .10 falls in this interval. We have

$$c_2 = .20 + 1.96\sqrt{\frac{.2(.8)}{100}}$$

$$= .20 + 1.96\sqrt{\frac{.4)}{10}} = .278$$

$$c_1 = .20 - 1.96\sqrt{\frac{.2(.8)}{100}}$$

$$= .20 - 1.96\frac{(.4)}{10} = .122$$

Since .10 is not in the interval (.122, .278), we can be more confident that we are not observing the placebo effect.

- The assumption made is that the normal approximation to the binomial distribution is reasonable here. In particular, we assume that  $\hat{p} \sim N\left(p, \frac{pq}{n}\right)$ . This assumption is appropriate since  $n\hat{p}\hat{q} = 100(.2)(.8) = 16 \ge 5$ .
- **6.43** The standard error  $=\frac{s}{\sqrt{n}} = \frac{12}{10} = 1.2$ .

**6.44** We can form a 95% confidence interval of the form  $(c_1, c_2)$  based on the t distribution where

$$c_{1} = \overline{d} + t_{n-1,025} \frac{s}{\sqrt{n}}$$

$$c_{2} = \overline{d} + t_{n-1,975} \frac{s}{\sqrt{n}}$$

Since 99 df is not in our t tables, we approximate  $t_{99,.975}$  using MINITAB as follows:

Thus,

$$c_1 = 5.3 - 1.9842(1.2) = 2.9$$
  
 $c_2 = 5.3 + 1.9842(1.2) = 7.7$ 

and we have the confidence interval (2.9, 7.7) .

- 6.45 Since the 95% confidence interval does not include 0, we can again conclude that the drug is effective using this measure of effectiveness.
- 6.46 The 95% confidence interval means that if a large number of samples of 100 hypertensive patients were selected and the preceding type of interval constructed, then approximately 95% of such intervals would contain the underlying mean difference in blood pressure upon using the drug.