

Manipulating Joint Distributions

Obtaining Marginal Density from Joint Density (continuous case)

$$f_W(w) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f_{W,X,Y,Z}(w, x, y, z) dx dy dz$$

$$f_{W,Y}(w, y) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f_{W,X,Y,Z}(w, x, y, z) dx dz$$

$$f_{X,Y,Z}(x, y, z) = \int_{-\infty}^{\infty} f_{W,X,Y,Z}(w, x, y, z) dw$$

Conditional Densities

$$f_{W|X,Y,Z}(w|x, y, z) = \frac{f_{W,X,Y,Z}(w, x, y, z)}{f_{X,Y,Z}(x, y, z)}$$

$$f_{X,Z|W,Y}(x, z|w, y) = \frac{f_{W,X,Y,Z}(w, x, y, z)}{f_{W,Y}(w, y)}$$

$$f_{X,Y,Z|W}(x, y, z|w) = \frac{f_{W,X,Y,Z}(w, x, y, z)}{f_W(w)}$$

Joint density as Product

$$\begin{aligned} & f_{W,X,Y,Z}(w, x, y, z) \\ &= f_W(w) f_{X|W}(x|w) f_{Y|W,X}(y|w, x) f_{Z|W,X,Y}(z|w, x, y) \end{aligned}$$

Introduction

$X_0, X_1, X_2, \dots$

is a random process.

(sequence of rv's defined on the same probability space).

Think of system evolving randomly over time:

X_i = state of system at time i .

$X_i \in S \quad \forall i$ (the state space).

For now assume S is discrete, and finite or countable.

Say $S = \{1, 2, \dots, m\}$

or $S = \{1, 2, 3, \dots\}$.

A Markov Process satisfies

$$(1) \quad P(X_{n+1} = i_{n+1} | X_n = i_n, \dots, X_0 = i_0)$$

$$= P(X_{n+1} = i_{n+1} | X_n = i_n)$$

for all $n \geq 0$

and all possible $i_0, i_1, \dots, i_{n+1} \in S$

(for which the conditioning event has probability > 0).

A Markov Process which has
stationary transition probabilities

(is time homogeneous)
satisfies

$$(2) \quad P(X_{n+1} = j | X_n = i) = P(X_1 = j | X_0 = i)$$

for all $n \geq 0, i, j \in S$.

A Markov Chain satisfies both
(1) and (2):

$$P(X_{n+1}=j | X_n=i, X_{n-1}=i_{n-1}, \dots, X_0=i_0) \\ = P(X_1=j | X_0=i) \equiv p_{ij}$$

for all ... (for which the conditioning event has prob > 0).

A Markov chain is described by its
transition prob. matrix $P = (p_{ij})$, and
initial distribution $a = (a_i)$

where $a_i \equiv P(X_0=i)$
for all $i \in S$.

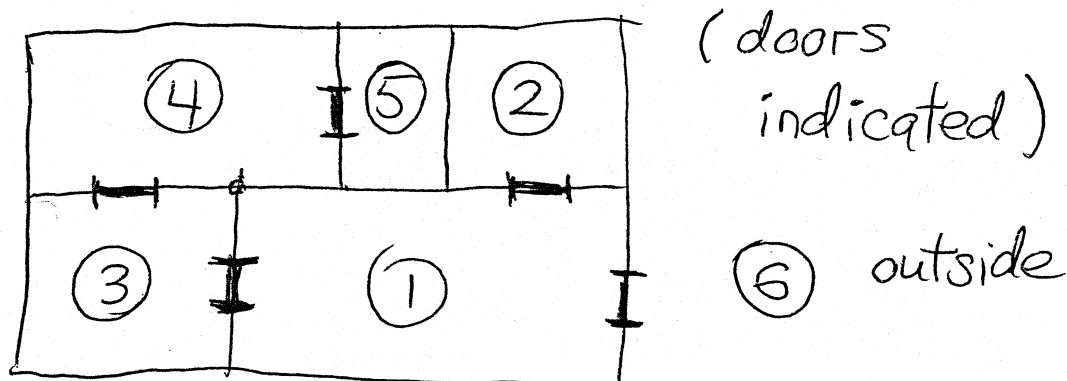
If S is infinite, these are
infinite-dimensional.

Clearly $p_{ij} \geq 0$, $a_i \geq 0$,

$$\sum_{j \in S} p_{ij} = 1 \quad \forall i \in S, \quad \sum_{i \in S} a_i = 1.$$

Example

Sleepwalker in house:



Starts in room 4 ($X_0 = 4$)

At each time $t = 1, 2, 3, \dots$

moves according to rule:

with prob $1/2$ stay in same room;
otherwise choose from available doors
at random (equally likely) and go
through.

Door to outside swings shut and locks.

Position of Sleepwalker forms a Markov
(location)

chain $X_0, X_1, X_2, \dots$

with state space $S = \{1, 2, \dots, 6\}$.

Initial distribution is

$$a = (a_1, \dots, a_6) = (0, 0, 0, 1, 0, 0).$$

Transition prob. matrix is

$$P = \begin{pmatrix} 1/2 & 1/6 & 1/6 & 0 & 0 & 1/6 \\ 1/2 & 1/2 & 0 & 0 & 0 & 0 \\ 1/4 & 0 & 1/2 & 1/4 & 0 & 0 \\ 0 & 0 & 1/4 & 1/2 & 1/4 & 0 \\ 0 & 0 & 0 & 1/2 & 1/2 & 0 \\ 0 & 0 & 0 & 0 & 0 & 1 \end{pmatrix}$$

Modified Sleepwalker Chain:

Eliminate state 6 (bolt the door).

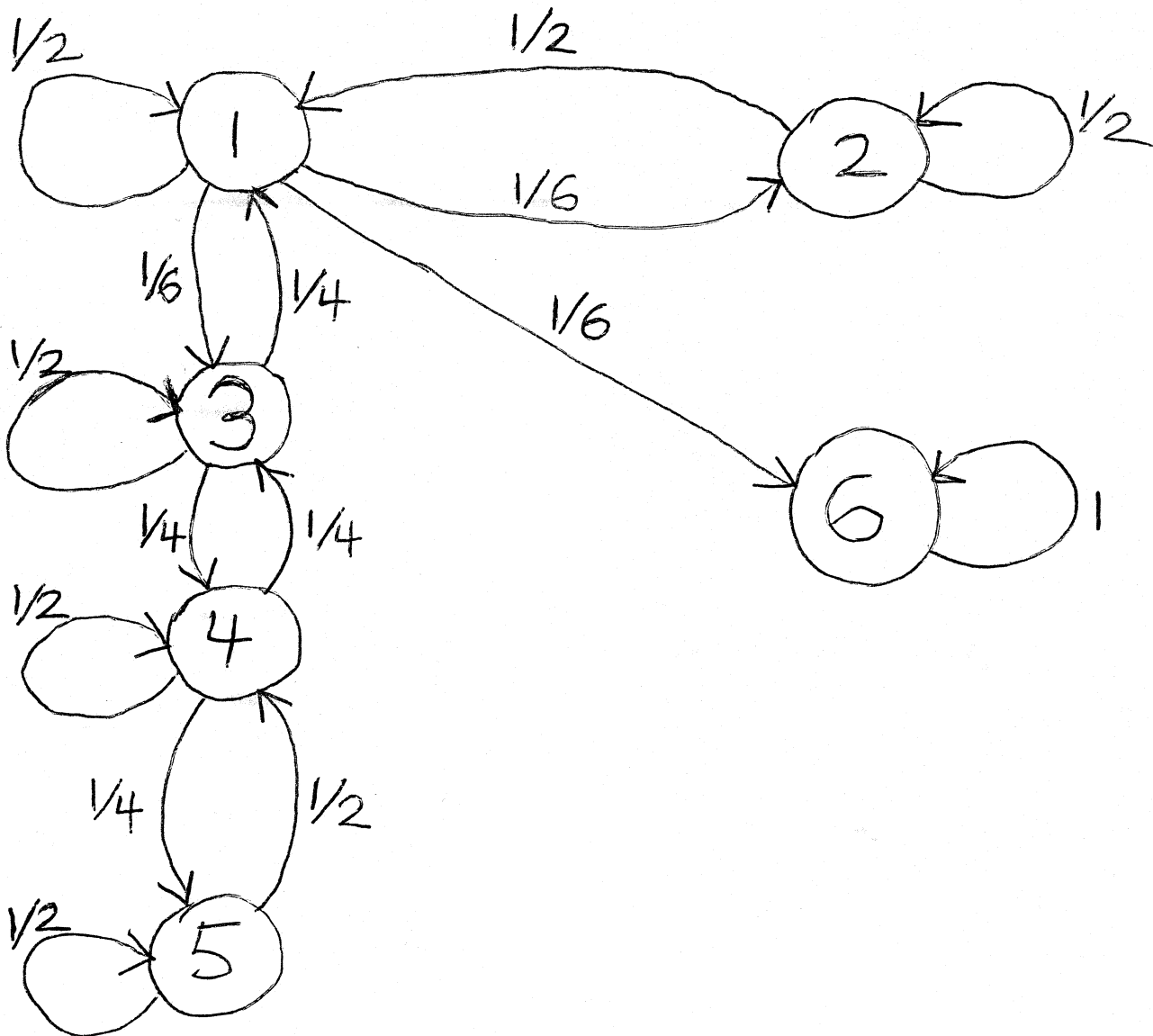
This chain has 5 states (rooms) and transition prob. matrix obtained by deleting the last row and column above, and changing the first row to read $(\frac{1}{2} \quad \frac{1}{4} \quad \frac{1}{4} \quad 0 \quad 0)$.

Picturing a transition matrix

Draw a node for each state $i \in S$.

Draw an arrow connecting i and j
if $p_{ij} > 0$.

Label arrow with value of p_{ij} .



Fifty Realizations of Sleepwalker Chain, Starting in Room 4

[illegible]

Fifty Realizations of Modified Sleepwalker Chain, Starting in Room 4

44311333345455445434344554444545554455544433331212221134543131111
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Comments

For finite state Markov chains, many properties are intuitively obvious.

For the original sleepwalker chain:

You will eventually exit the house
(enter the absorbing state).

All rooms (inside) will be visited
only finitely many times (perhaps
zero times). (†)

For the modified sleepwalker chain
(bolting the door to the outside):

Starting from any room, you are
guaranteed to (eventually) return.

All rooms will be visited infinitely
often.

For infinite state Markov chains, things are
not so clear.

For a house with infinitely many rooms,
all of the earlier statements could be false!
(statement † is still true, but this is
a subtle point.)

$$P(X_0=i, X_1=j, X_2=k, X_3=l) \\ = a_i p_{ij} p_{jk} p_{kl}$$

In general,

$$P(X_0=i_0, X_1=i_1, \dots, X_n=i_n) \\ = a_{i_0} p_{i_0 i_1} p_{i_1 i_2} \dots p_{i_{n-1} i_n}$$

This is essentially a consequence of
 $P(A \cap B \cap C \cap D) = P(A) P(B|A) P(C|A \cap B) \nearrow$
 $P(D|A \cap B \cap C)$

or of

$$f_{W,X,Y,Z}(w,x,y,z) = f_W(w) f_{X|W}(x|w) \nearrow \\ f_{Y|W,X}(y|w,x) f_{Z|W,X,Y}(z|w,x,y)$$

Fact:

$$P(X_1=i_1, \dots, X_n=i_n | X_0=i_0) \quad (*)$$

$$= P_{i_0 i_1} P_{i_1 i_2} \dots P_{i_{n-1} i_n}$$

$$\text{if } P(X_0=i_0) > 0.$$

Notation:

$P_i(\cdot)$ = probability measure of MC
started from state i ,
that is, with initial distribution

$$a = \delta_i = (\delta_{ij} : j \in S)$$

$$\text{where } \delta_{ij} = \begin{cases} 1 & \text{if } j=i \\ 0 & \text{if } j \neq i \end{cases}$$

Fact restated:

$$P_{i_0}(X_1=i_1, \dots, X_n=i_n) = P_{i_0 i_1} P_{i_1 i_2} \dots P_{i_{n-1} i_n}$$

Proof: $(*) = P(X_0=i_0, \dots, X_n=i_n) / P(X_0=i_0)$

$$= a_{i_0} P_{i_0 i_1} \dots P_{i_{n-1} i_n} / a_{i_0} \quad \text{QED}$$

Higher Order Transition Probabilities

Notation: $P^n = (p_{ij}^{(n)} : i \in S, j \in S)$.

Facts: (about MC with t.p.m. P)

$$P(X_n = k) = \sum_{j \in S} a_j p_{jk}^{(n)} = (aP^n)_k \quad (1)$$

(aP^n gives the distn. of X_n .)

$$P_i(X_n = k) = p_{ik}^{(n)} \quad (2)$$

(P^n gives the n -step transition probabilities.)

$$P_{i_0}(X_{t_1} = i_1, \dots, X_{t_n} = i_n) \quad (0 < t_1 < t_2 < \dots < t_n)$$

$$= p_{i_0 i_1}^{(t_1)} p_{i_1 i_2}^{(t_2 - t_1)} \dots p_{i_{n-1} i_n}^{(t_n - t_{n-1})} \quad (3)$$

Note: P^k is a "stochastic" matrix (for $k \geq 1$)

↳ (nonnegative elements,
rows sum to 1.)

and $X_0, X_k, X_{2k}, X_{3k}, \dots$ is a MC
with t.p.m. P^k .

$$P(X_n = k) = (aP^n)_k. \text{ Why?}$$

Follows from "Law of Total Probability" and induction.

$$\begin{aligned} P(X_{n+1} = k) &= \sum_{j \in S} P(X_{n+1} = k | X_n = j) P(X_n = j) \\ &= \sum_{j \in S} P(X_n = j) p_{jk} \end{aligned}$$

But this exactly says

$$\pi^{n+1} = \pi^n P$$

where π^{n+1}, π^n are row vectors with entries:

$$\pi_k^n = P(X_n = k), \quad \pi_k^{n+1} = P(X_{n+1} = k).$$

Using this, starting from $\pi^0 = a$, gives

$$\pi^1 = aP$$

$$\pi^2 = (aP)P = aP^2$$

$$\pi^3 = (aP^2)P = aP^3$$

etc.

The Sleepwalker Transition Probability Matrix

P1

	[,1]	[,2]	[,3]	[,4]	[,5]	[,6]
[1,]	0.50	0.16667	0.16667	0.00	0.00	0.16667
[2,]	0.50	0.50000	0.00000	0.00	0.00	0.00000
[3,]	0.25	0.00000	0.50000	0.25	0.00	0.00000
[4,]	0.00	0.00000	0.25000	0.50	0.25	0.00000
[5,]	0.00	0.00000	0.00000	0.50	0.50	0.00000
[6,]	0.00	0.00000	0.00000	0.00	0.00	1.00000

P

$$P_{ij} = p_{ij}^{(1)} = P_i(X_1=j)$$

P2 <- P1 %*% P1

P2

	[,1]	[,2]	[,3]	[,4]	[,5]	[,6]
[1,]	0.3750	0.166667	0.166667	0.041667	0.0000	0.250000
[2,]	0.5000	0.333333	0.083333	0.000000	0.0000	0.083333
[3,]	0.2500	0.041667	0.354167	0.250000	0.0625	0.041667
[4,]	0.0625	0.000000	0.250000	0.437500	0.2500	0.000000
[5,]	0.0000	0.000000	0.125000	0.500000	0.3750	0.000000
[6,]	0.0000	0.000000	0.000000	0.000000	0.0000	1.000000

P²

$$(P^2)_{ij} = p_{ij}^{(2)} = P_i(X_2=j)$$

P4 <- P2 %*% P2

P4

	[,1]	[,2]	[,3]	[,4]	[,5]	[,6]
[1,]	0.26823	0.125000	0.14583	0.075521	0.020833	0.364583
[2,]	0.37500	0.1979167	0.14062	0.041667	0.0052083	0.239583
[3,]	0.21875	0.0703125	0.24089	0.239583	0.1080729	0.1223958
[4,]	0.11328	0.0208333	0.23958	0.381510	0.2187500	0.0260417
[5,]	0.06250	0.0052083	0.21615	0.437500	0.2734375	0.0052083
[6,]	0.00000	0.0000000	0.00000	0.000000	0.0000000	1.0000000

P⁴

$$(P^4)_{ij} = p_{ij}^{(4)} = P_i(X_4=j)$$

P32

	[,1]	[,2]	[,3]	[,4]	[,5]	[,6]
[1,]	0.038470	0.013874	0.043501	0.055191	0.029626	0.81934
[2,]	0.041623	0.015055	0.046831	0.059252	0.031776	0.80546
[3,]	0.065251	0.023415	0.074426	0.094877	0.051011	0.69102
[4,]	0.082787	0.029626	0.094877	0.121257	0.065251	0.60620
[5,]	0.088878	0.031776	0.102022	0.130502	0.070246	0.57658
[6,]	0.000000	0.000000	0.000000	0.000000	0.000000	1.00000

P^{32}

$$(P^{32})_{ij} = P_{ij}^{(32)} = P_i(X_{32}=j)$$

> P128

	[,1]	[,2]	[,3]	[,4]	[,5]	[,6]
[1,]	0.0012253	0.00043931	0.0013997	0.0017858	0.00096039	0.99419
[2,]	0.0013179	0.00047252	0.0015055	0.0019208	0.00103300	0.99375
[3,]	0.0020996	0.00075276	0.0023984	0.0030599	0.00164565	0.99004
[4,]	0.0026787	0.00096039	0.0030599	0.0039039	0.00209955	0.98730
[5,]	0.0028812	0.00103300	0.0032913	0.0041991	0.00225829	0.98634
[6,]	0.0000000	0.00000000	0.0000000	0.0000000	0.00000000	1.00000

P^{128}

$$(P^{128})_{ij} = P_{ij}^{(128)} = P_i(X_{128}=j)$$

$$P_4(X_1=5, X_2=5, X_3=4, X_4=3)=$$

$$P1[4,5]*P1[5,5]*P1[5,4]*P1[4,3] = p_{45} p_{55} p_{54} p_{43}$$

0.015625

P4[4,3]

0.23958

$$P_4(X_4=3) = p_{43}^{(4)}$$

P1[5,4]*P1[4,3]*P1[3,1]*P1[1,6]

0.0052083

$$P_5(X_1=4, X_2=3, X_3=1, X_4=6)$$

P4[5,6]

0.0052083

$$P_5(X_4=6)$$

P4[5,3]*P4[3,2]*P2[2,1]*P1[1,3]

0.0012665

$$= p_{53}^{(4)} p_{32}^{(4)} p_{21}^{(2)} p_{13}^{(1)}$$

$$= P_5(X_4=3, X_8=2, X_{10}=1, X_{11}=3)$$

Communication

Let i and j be states.

Definition:

$i \rightarrow j$ means $p_{ij}^{(n)} > 0$ for some $n \geq 0$.

$i \rightarrow j$ is read as “ j is accessible from i ”.

Note: $P^0 = I$ and $p_{ij}^{(0)} = \begin{cases} 1 & \text{if } i = j \\ 0 & \text{if } i \neq j \end{cases}$.

Since $n = 0$ is allowed in the definition of “ $\rightarrow$ ” and $p_{ii}^{(0)} = 1 > 0$, it is always true that $i \rightarrow i$.

However, $i \rightarrow i$ does **not** imply that, starting from i , it is possible to return to i at a later time.

Definition:

$i \leftrightarrow j$ means $i \rightarrow j$ and $j \rightarrow i$.

$i \leftrightarrow j$ is read as “ i and j communicate”.

Communication is an equivalence relation:

- $i \leftrightarrow i$ (reflexive).
- If $i \leftrightarrow j$, then $j \leftrightarrow i$. (symmetric)
- If $i \leftrightarrow j$ and $j \leftrightarrow k$, then $i \leftrightarrow k$. (transitive)

Therefore, the state space S is divided (partitioned) into disjoint equivalence classes:

$$S = \bigcup_{\alpha} C_{\alpha}$$

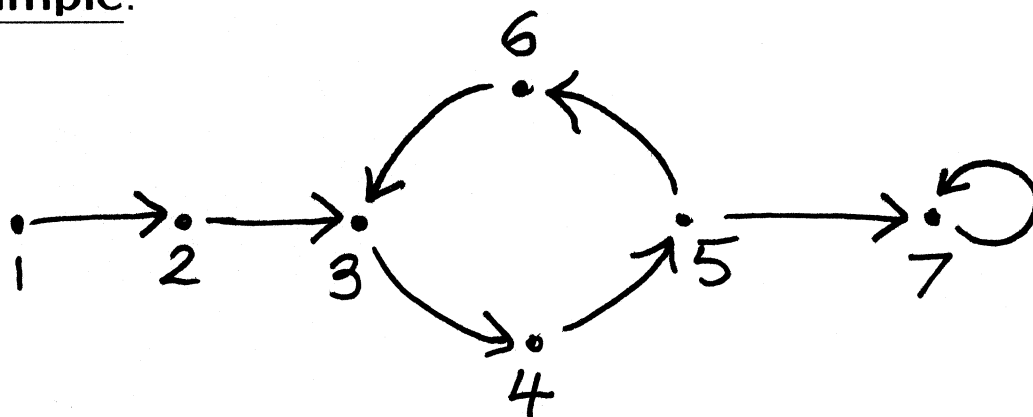
where $i \leftrightarrow j$ if and only if i and j belong to the same class C_{α} .

The states in a class C_α communicate with each other, but **not** with any states outside C_α .

The classes C_α are called the *communication* classes.

If the state space S consists of a single communication class (that is, all states communicate), then the Markov chain is said to be irreducible.

Example:



Arrows indicate positive transition probabilities.

The communication classes are:

$\{1\}, \{2\}, \{3, 4, 5, 6\}, \{7\}$

Period

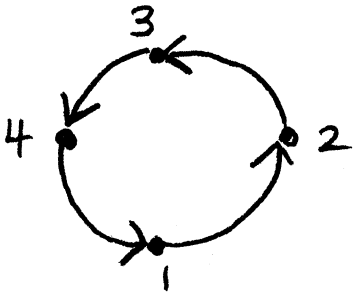
For $i \in S$,

$$d(i) = \text{gcd} \{ n : p_{ii}^{(n)} > 0 \}.$$

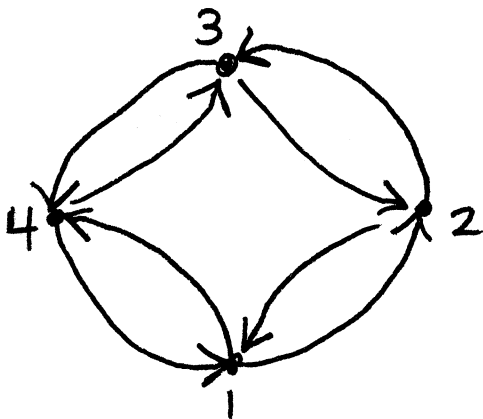
period of i greatest common divisor The set of possible return times

If $d(i) = 1$, we say i is aperiodic.

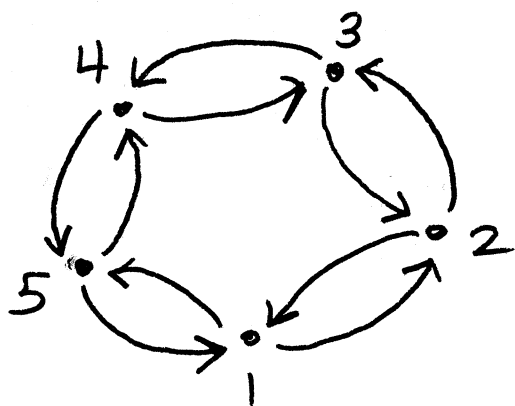
Examples



$$d(i) = 4 \quad \forall i.$$
$$p_{ii}^{(4n)} = 1 \quad \text{for all } i \text{ and } n.$$
$$\{n : p_{ii}^{(n)} > 0\} = \{0, 4, 8, \dots\}$$



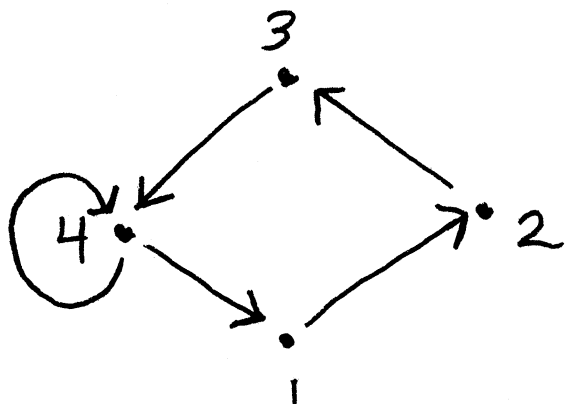
$$d(i) = 2 \quad \forall i$$
$$\{n : p_{ii}^{(n)} > 0\} = \{0, 2, 4, \dots\}$$



$$d(i) = 1 \quad \forall i$$

$$\{n: p_{ii}^{(n)} > 0\}$$

$$= \{0, 2, 4, 5, 6, 7, \dots\}$$

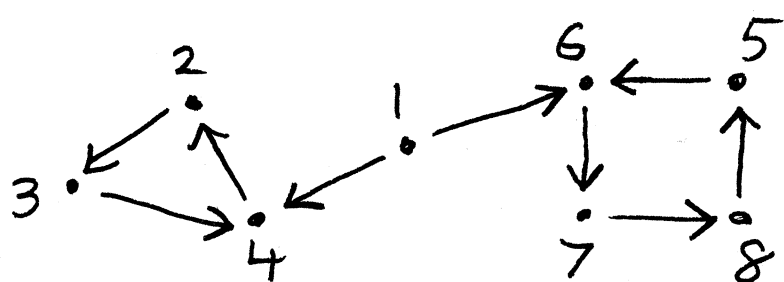


$$d(i) = 1 \quad \forall i$$

$$\{n: p_{44}^{(n)} > 0\}$$

$$= \{0, 1, 2, 3, \dots\}$$

$$\{n: p_{11}^{(n)} > 0\} = \{0, 4, 5, 6, \dots\}$$



$$d(1) = \infty.$$

$$d(i) = 3$$

for $i \in \{2, 3, 4\}.$

$$d(i) = 4$$

for $i \in \{5, 6, 7, 8\}$

Period is a class property:

If $i \leftrightarrow j$, then $d(i) = d(j).$