## Ph.D. Qualifying Exam Monday–Tuesday, January 4–5, 2016

Put your solution to each problem on a separate sheet of paper.

**Problem 1.** (5106) Find the maximum likelihood estimate of  $\theta$  where  $\theta$  is a parameter in the multinomial distribution:

$$(x_1, x_2, x_3, x_4) \sim M(n; 0.4(1+\theta), 0.3\theta, 0.6(1-2\theta), 0.5\theta)$$

- (a) Choose a variable for the missing data and use the EM algorithm for iteratively estimating  $\theta$ . Let  $\theta^{(m)}$  be the current value of the unknown. Derive the mathematical formula to update for  $\theta^{(m+1)}$ .
- (b) Let  $L(\theta)$  denote the likelihood with the parameter  $\theta$ . Prove that

$$L(\theta^{(m+1)}) \ge L(\theta^{(m)}).$$

**Problem 2.** (5106) Let  $x = (x_1, \dots, x_n)$  be a given binary, Markovian sequence. In particular,

$$P(x_1 = 0) = 0.5, P(x_1 = 1) = 0.5,$$

and

$$P(x_i = x_{i-1}) = p, \ P(x_i = 1 - x_{i-1}) = 1 - p,$$

 $i = 2, \dots, n$ . Let  $y = (y_1, \dots, y_n)$  be a noisy observation of x. That is,

 $y_i = x_i + e_i,$ 

with  $e_i \sim N(0, \sigma^2), \, i = 1, \cdots, n.$ 

Assuming p and  $\sigma^2$  known, we can use the Maximum A Posteriori method to reconstruct x from y in the following form:

$$\hat{x} = \operatorname{argmax}_{\{x_i\}} \sum_{i=1}^n \left(-\frac{(y_i - x_i)^2}{2\sigma^2}\right) + \sum_{i=2}^n \log(1_{x_i = x_{i-1}}p + 1_{x_i \neq x_{i-1}}(1-p))$$

Write out a pseudocode for a Dynamic Programming procedure to compute  $\hat{x}$  in the computational order of n.

**Problem 3.** (5166) Let  $\{X_1, X_2, \ldots, X_n\}$  be a sample generated from the model:

$$X_i = \epsilon_i + \theta \epsilon_{i-1},$$

where the  $\epsilon_i$ 's are independent and identically distributed random variables with mean zero and variance  $\sigma^2$ . Let  $\bar{X} = \frac{1}{n} \sum_{i=1}^n X_i$ .

(a) Find the variance of  $\bar{X}$ .

(b) Let 
$$s^2 = \frac{\sum_{i=1}^n (X_i - \bar{X})^2}{n-1}$$
. Find  $E(s^2)$ .

(c) Compare Var $(\bar{X})$  with  $E(s^2/n)$  for  $\theta = 0, \theta > 0$ , and  $\theta < 0$ .

## **Problem 4.** (5166)

Consider a  $2^{5-2}$  fractional factorial design with generators 4=12 and 5=13. After analyzing the results from the design the experimenter decided to perform a second  $2^{5-2}$  design exactly the same as the first one but switching the signs of column 1 from "+" to"-" (a fold-over design).

- (a) What are the advantages and disadvantages of a fractional factorial design compared with a full factorial design?
- (b) What are the defining relation and design resolution of the first design? List the confounding patterns of all the two-factor interactions in the first design.
- (c) What are the defining relation and design resolution of the combined design? List the confounding patterns of all the two-factor interactions in the combined design.

**Problem 5.** (5167) In a one-way layout random-effects experiment, consider the following linear model:

$$Y_{ti} = \mu + \tau_t + \epsilon_{ti}, \quad t = 1, \dots, k; \quad i = 1, \dots, n_t,$$

where  $\mu$  is the overall mean,  $\{\tau_t, t = 1, \ldots, k\}$  are iid  $N(0, \sigma_{\tau}^2)$ ,  $\{\epsilon_{ti}, t = 1, \ldots, k; i = 1, \ldots, n_t\}$  are iid  $N(0, \sigma_{\epsilon}^2)$ , and the  $\tau_t$ 's and  $\epsilon_{ti}$ 's are independent.

- (a) Let  $N = \sum_{t=1}^{k} n_t$ ,  $\bar{Y}_{t.} = \frac{1}{n_t} \sum_{i=1}^{n_t} Y_{ti}$ , and  $\bar{Y} = \frac{1}{N} \sum_{t=1}^{k} \sum_{i=1}^{n_t} Y_{ti}$ . Calculate the expected values and variances of  $\bar{Y}_{t.}$  and  $\bar{Y}$ .
- (b) Find  $\operatorname{Cov}(\bar{Y}_{t}, \bar{Y})$ .
- (c) Let  $S_T = \sum_{t=1}^k n_t (\bar{Y}_{t.} \bar{Y})^2$ ,  $S_R = \sum_{t=1}^k \sum_{i=1}^{n_t} (Y_{ti} \bar{Y}_{t.})^2$ , and  $S_D = \sum_{t=1}^k \sum_{i=1}^{n_t} (Y_{ti} - \bar{Y})^2$ . Show that  $S_D = S_T + S_R$ . Find the expected values of  $S_T$  and  $S_R$ .

**Problem 6.** (5167) Consider the linear regression model with p predictors:

$$Y = X\boldsymbol{\beta} + \boldsymbol{e},$$

where  $Y = (y_1, \ldots, y_n)'$ ,  $\boldsymbol{e} = (e_1, \ldots, e_n)'$  with  $e_1, e_2, \ldots, e_n$  iid  $N(0, \sigma^2)$ ,  $\boldsymbol{\beta} = (\beta_0, \beta_1, \ldots, \beta_p)'$ , and X is an  $n \times (p+1)$  full-rank matrix. Let  $\hat{\boldsymbol{\beta}}$  be the least squares estimate of  $\boldsymbol{\beta}$ .

- (a) Define  $\hat{Y} = X\hat{\beta}$  and  $\hat{e} = Y \hat{Y}$ . Find the distribution of  $\hat{e}$ .
- (b) Let  $\mathbf{1} = (1, 1, \dots, 1)'$ , a column of ones with length n. Define

$$SS_{req} = (\hat{Y} - \bar{Y}\mathbf{1})'(\hat{Y} - \bar{Y}\mathbf{1}).$$

Find the distribution of  $SS_{reg}$  under  $H_0: \boldsymbol{\beta} = 0$ .

(c) Show that the statistic

$$F = \frac{SS_{reg}/p}{\hat{e}'\hat{e}/(n-p-1)}$$

has an *F*-distribution under  $H_0: \beta = 0$ .

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**Problem 7.** (5326) An urn contains R red balls and G green balls. Ed draws out K balls one by one, randomly and with**OUT** replacement. Whenever Ed draws two red balls in a row he wins a dollar. To be precise, Ed receives a dollar after the  $i^{\text{th}}$  draw if he draws red balls on draws i - 1 and i. Note that under these rules, if Ed draw 3 red balls in a row he receives 2 dollars; if he draws 4 red balls in a row he receives 3 dollars, etc. Let X be the amount of Ed's total winnings.

- (a) Find EX.
- (b) Find  $\operatorname{Var}(X)$ .

**Problem 8.** (5326) Consider the Gambler's Ruin problem: a gambler is betting on the tosses of a **fair** coin, winning \$1 when it comes up heads and losing \$1 when it comes up tails. The gambler has an initial fortune of z dollars and a goal of g dollars, i.e., starting with z dollars, he continues playing until he reaches g dollars or goes broke (loses all his money).

- (a) Let ψ(z) denote the probability of reaching the goal g as a function of the initial fortune z. Use the Law of Total Probability to find an equation that ψ(z) must satisfy. Then find ψ(z) by obtaining a solution of this equation which satisfies ψ(0) = 0 and ψ(g) = 1.
- (b) Given that the gambler reached the goal, what is the probability that his first six tosses had a total of four heads and two tails? (Use your answer to part (a). To avoid complications, assume  $6 \le z \le g 6$  so that the gambler makes at least six tosses.)

**Problem 9.** (5327) Let  $X_i$  be independently distributed as Poisson $(\beta N_i)$  for i = 1, 2, ..., n where  $\beta$  is an unknown positive number and  $N_1, N_2, ..., N_n$  are known positive constants.

- (a) Find the maximum likelihood estimate (MLE) of  $\beta$ . Call it  $T_1$ .
- (b) Show that  $T_1$  is unbiased for  $\beta$ .
- (c) Observe that  $T_2 = \frac{1}{n} \sum_{i=1}^{n} \frac{X_i}{N_i}$  is also an unbiased estimator of  $\beta$ . Which one of  $T_1$  and  $T_2$  has smaller variance? Prove your claim.

**Problem 10.** (5327) Let  $X_1, X_2, \ldots, X_n$  be i.i.d from a distribution with pdf given by

$$f(x \mid \theta) = \begin{cases} \theta^{-c} c x^{c-1} e^{-(x/\theta)^c}, \text{ if } x > 0\\ 0, \text{ otherwise,} \end{cases}$$

where c > 0 is a known constant and  $\theta > 0$  is the unknown parameter.

- (a) Find a complete sufficient statistic for  $\theta$ .
- (b) Find the minimum variance unbiased estimator (MVUE) for  $\theta$ . (Hint: Find an unbiased statistic by taking a suitable power of the complete sufficient statistic.)
- (c) Find the most powerful (MP) test of size  $\alpha$  for testing

$$H_0: \theta = \theta_0, \quad \text{vs} \quad H_1: \theta = \theta_1,$$

where  $0 < \theta_0 < \theta_1$  are known constants.

**Problem 11.** (6346) Suppose  $(\Omega, \mathcal{F}_0, \mu)$  is a probability space,  $\mathcal{F}$  is a  $\sigma$ -field with  $\mathcal{F} \subset \mathcal{F}_0, X : \Omega \to \mathbb{R}$  is measurable with respect to  $\mathcal{F}_0$ , and  $\mathbb{E}|X| < \infty$ .

- (a) Define conditional expectation,  $\mathbb{E}(X|\mathcal{F})$ .
- (b) Let  $Y = I_A$  and  $X = I_B$  for some  $A, B \in \mathcal{F}_0$  with  $0 < \mu(A) < 1$ . Let  $\mathcal{G} = \sigma(Y)$ . Find  $\mathbb{E}(X|\mathcal{G})$ .
- (c) What happens if A and B are the same set?

**Problem 12.** (6346) Let B(t) be standard Brownian motion.

- (a) Show that the joint density of  $B(t_1) = X_1, B(t_2) = X_2, \ldots, B(t_n) = X_n$  where  $0 < t_1 < \cdots < t_n$  is a product of univariate normal densities.
- (b) Find  $E[B(t)|B(s)], 0 \le s \le t$ .
- (c) Define X(t) = B(1/t), t > 0, and X(0) = 0. Describe the process X(t) in terms of its mean and covariance. Is X(t) Brownian motion? Is it a Gaussian process?