

Final Project: Reversible Jump Markov Chain
Monte Carlo Algorithm for Model Selection in
Linear Regression
Anuj Srivastava

Jaime Frade

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Chapter 1

Introduction

1.1 Problem Statement:

The objective of this assignment is to solve for the regression solving for the regression coefficients in a standard linear model, yet the selection of predictors is not known. In other words, we are given a large number, say $m = 10$, of predictors and we have to select an appropriate subset to obtain the optimal model.

This model selection problem is the main focus of this project. We will restrict to a smaller problem where the optimal subset is simply the first n predictors, yet n is unknown. This reduces the possible number of models from the power set of a finite set with $m, 2^m$ elements, to only m possibilities.

Chapter 2

Methodology

2.1 Preliminaries

In this assignment we seek coefficients for the model

$$y = \sum_{i=1}^n x_i b_i + \epsilon \quad (2.1)$$

where $n < m$, x_i 's are the predictors, y is the response variable, and ϵ is the measurement noise. In this problem, given k independent measurements, denoted by \mathbf{y} , \mathbf{X} , and ϵ . To solve this problem, we will formulate a Bayesian solution to the joint estimation of $\{n, b_1, \dots, b_n\}$.

2.2 Definitions

To setup a Bayesian formulation, define a joint posterior density of the type:

$$f(n, \mathbf{b}_n | \mathbf{y}) \propto f(\mathbf{y}, n | \mathbf{b}_n) f(\mathbf{b}_n | n) f(n)$$

Let the notation $X_n = X(:, 1 : n)$ and $\mathbf{b}_n = \{b_1, \dots, b_n\}$. In this project, we will use the following terms

The likelihood function is given by

$$f(\mathbf{y}, n | \mathbf{b}_n) = \left(\frac{1}{\sqrt{2\pi\sigma_0^2}} \right)^k e^{-\frac{1}{2\sigma_0^2} \|\mathbf{y} - \mathbf{X}_n \mathbf{b}_n\|^2} \quad (2.2)$$

The prior on \mathbf{b}_n given n is:

$$f(\mathbf{b}_n | n) = \left(\frac{1}{\sqrt{2\pi\sigma_p^2}} \right)^k e^{-\frac{1}{2\sigma_p^2} \|\mathbf{b}_n - \mu_b\|^2} \quad (2.3)$$

The prior on n is uniform:

$$f(n) = \frac{1}{m} \quad (2.4)$$

2.3 Sampling from the Posterior

The method of the project will involve using the Reversible Jump Monte Carlo sampling (RJMCMC) algorithm for sampling from the posterior.

Let (n, b_n) be the current samples from the posterior. The following is the algorithm used to apply RJMCMC:

- (a) Select a candidate number, n^* from the probability $f(n)$.
- (b) If $n^* \leq n$, generate a random vector $u \sim N(0, \sigma_r I_{n^*})$. The candidate coefficient vector is given by:

$$b_{n^*} = \begin{bmatrix} \mathbf{b}_n \\ 0 \end{bmatrix} + \begin{bmatrix} \mathbf{u}_1 \\ \mathbf{u}_2 \end{bmatrix}, \quad \mathbf{u} = \begin{bmatrix} \mathbf{u}_1 \\ \mathbf{u}_2 \end{bmatrix}$$

Compute the likelihoods:

$$\begin{aligned} h_1(\mathbf{u}) &= \left(\frac{1}{\sqrt{2\pi\sigma_r^2}} \right)^{n^*} e^{-\frac{1}{2\sigma_r^2} \|\mathbf{u}\|^2} \\ h_2(\mathbf{u}_1) &= \left(\frac{1}{\sqrt{2\pi\sigma_r^2}} \right)^n e^{-\frac{1}{2\sigma_r^2} \|\mathbf{u}_1\|^2} \end{aligned}$$

- (c) If $n^* < n$, generate a random vector $u_1 \sim N(0, \sigma_r I_{n^*})$. The candidate coefficient vector is given by:

$$b_{n^*} = \mathbf{b}_n^1 + \mathbf{u}_1, \quad \mathbf{b}_n = \begin{bmatrix} \mathbf{b}_n^1 \\ \mathbf{b}_n^2 \end{bmatrix}$$

and form $\mathbf{u} = \begin{bmatrix} \mathbf{u}_1 \\ \mathbf{b}_n^2 \end{bmatrix}$

Compute the likelihoods:

$$\begin{aligned} h_2(\mathbf{u}) &= \left(\frac{1}{\sqrt{2\pi\sigma_r^2}} \right)^n e^{-\frac{1}{2\sigma_r^2} \|\mathbf{u}\|^2} \\ h_1(\mathbf{u}_1) &= \left(\frac{1}{\sqrt{2\pi\sigma_r^2}} \right)^{n^*} e^{-\frac{1}{2\sigma_r^2} \|\mathbf{u}_1\|^2} \end{aligned}$$

- (d) Compute the acceptance-rejection function:

$$\begin{aligned} \rho &= \min \left(1, \frac{f(n^*, \mathbf{b}_{n^*} | \mathbf{y}) h_2}{f(n, \mathbf{b}_n | \mathbf{y}) h_1} \right) = \min \left(1, \frac{f(\mathbf{y} | n^*, \mathbf{b}_{n^*}) f(\mathbf{b}_{n^*} | n^*) h_2}{f(\mathbf{y} | n, \mathbf{b}_n) f(\mathbf{b}_n | n) h_1} \right) \\ &= \min \left(1, \frac{e^{(E_1 - E_2)} (2\pi\sigma_p^2)^{\frac{n-n^*}{2}} e^{-\frac{1}{2\sigma_p^2} (\|\mathbf{b}_{n^*} - \mu_{\mathbf{b}}\|^2 - \|\mathbf{b}_n - \mu_{\mathbf{b}}\|^2)} h_2}{h_1} \right) \end{aligned}$$

Where $E_1 = \frac{1}{2\sigma_0^2} \|\mathbf{y} - \mathbf{X}_{n^*} \mathbf{b}_{n^*}\|^2$ and $E_2 = \frac{1}{2\sigma_0^2} \|\mathbf{y} - \mathbf{X}_n \mathbf{b}_n\|^2$

(e) If $U \sim U[0, 1]$ is less than ρ then set $(n, \mathbf{b}_n) = (n^*, \mathbf{b}_{n^*})$. Else, return to Step (a).

Chapter 3

Matlab Code

```
clear all
clc
N=100000;
m=10;
k=10;
sigma0=.2;
sigmap=.3;
sigmar=.2;

%initial conditions
n0=ceil(rand*m);
mub=ones(1,n0)*2;
b=normrnd(2,sigmap,n0,1);
X=5*randn(k,m);
y=X(:,1:n0)*b+sigma0*randn(k,1);
tic
for i=1:N
    nstar=ceil(rand*m);
    mustar=ones(1,nstar)*2;
    if (nstar<n0)
        u1=normrnd(0,sigma0,nstar,1);
        b1=b(1:length(u1));
        b2=b(length(u1)+1:end);
        bnstar=b1+u1;
        u=[u1;b2];
        Unorm = sum(u(:).^2);
        h2 = ((1/sqrt(2*pi*sigmar^2))^n0)*exp((-1/(2*sigmar^2))*(Unorm));
        U1norm = sum(u1(:).^2);
        h1 = ((1/sqrt(2*pi*sigmar^2))^nstar)*exp((-1/(2*sigmar^2))*(U1norm));
    else
        u=normrnd(0,sigma0,nstar,1);
        u1=u(1:n0);
        u2=u(n0+1:end);
        bn=[b;zeros(length(u2),1)];
```



```

        bnstar=bn+u;
        Unorm = sum(u(:).^2);
        h1 = ((1/sqrt(2*pi*sigmar^2))^nstar)*exp((-1/(2*sigmar^2))*(Unorm));
        U1norm = sum(u1(:).^2);
        h2 = ((1/sqrt(2*pi*sigmar^2))^n0)*exp((-1/(2*sigmar^2))*(U1norm));
    end
    Xstar=X(1:nstar,:);
    Xn=X(1:n0,:);
    y1=y'-bnstar'*Xstar;
    y2=y'-b'*Xn;
    y1 = sum(y1(:).^2);
    y2 = sum(y2(:).^2);
    E1=y1/(2*sigma0^2);
    E2=y2/(2*sigma0^2);
    E3 = bnstar-mustar';
    E3 = sum(E3(:).^2);
    E4 = b-mub';
    E4 = sum(E4(:).^2)';
    rho1=(exp(-(E1-E2)))*((2*pi*sigmap^2)^((n0-nstar)/2))*(exp((-1/(2*sigmap^2))*(E3-E4)))
    rho=min(rho1,1);

    U=rand;
    if(U<rho)
        n(i+1)=nstar;
    else
        n(i+1)=n0;
    end
end
toc
n;
figure(1)
hist(n)
h=hist(n);
idx=find(h==max(h))
b

```

Chapter 4

Results

4.1 Experimental Results

Simulated a dataset with the following code below. For this data (y, X) implement the RJMCMC algorithm (with $\sigma_r = 0.2$) to a sample and using $N = 100,000$ samples from the posterior.

```
m=10;
k=10;
sigma0=.2;
sigmap=.3;
sigmar=.2;
n0=ceil(rand*m);
mub=ones(1,n0)*2;
b=normrnd(2,sigmap,n0,1);
X=5*randn(k,m);
y=X(:,1:n0)*b+sigma0*randn(k,1);
```

4.1.1 Histogram plots of several different runs

Show a histogram of n values visited by Markov chain. This estimates the posterior probability, $f(n|\mathbf{y})$.

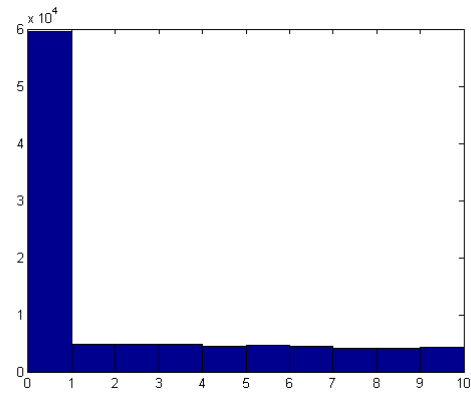


Figure 4.1: Histogram: $n=1$

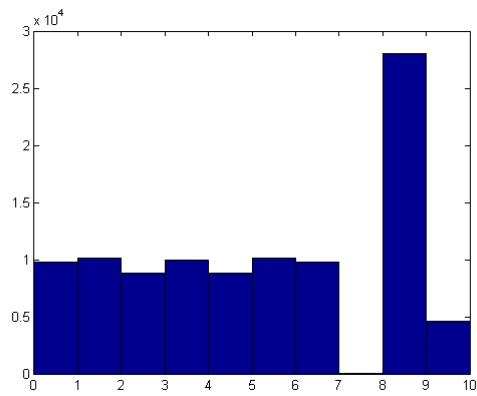


Figure 4.2: Histogram: $n=9$

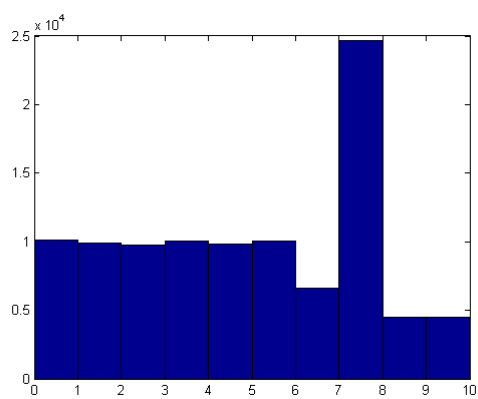


Figure 4.3: Histogram: $n=8$

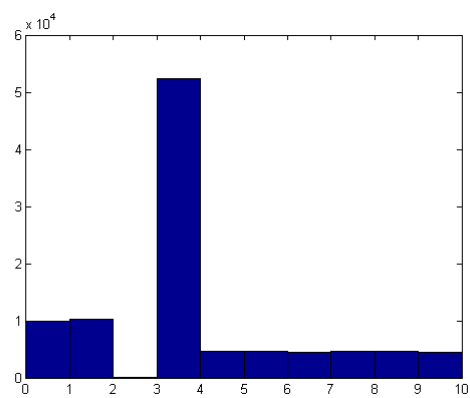


Figure 4.4: Histogram: $n=4$

4.1.2 Computation Times

| | |
|---------|---------------------|
| $n = 1$ | 481.608973 seconds |
| $n = 9$ | 668.774289 seconds. |
| $n = 8$ | 651.934132 seconds. |
| $n = 4$ | 533.207485 seconds. |

4.1.3 Table of Conditional Mean Estimates of \mathbf{b}_n

For the n value that is visited most often, find the conditional mean estimate of \mathbf{b}_n under the density, $f(\mathbf{b}_n|\mathbf{y})$

| n | b | n | b |
|-----|--------|-----|--------|
| 1 | 1.6652 | 9 | 2.4377 |
| | | | 1.9074 |
| | | | 1.8166 |
| | | | 1.9085 |
| | | | 2.1502 |
| | | | 2.2462 |
| | | | 1.8852 |
| | | | 2.1858 |
| | | | 2.1320 |
| n | b | n | b |
| 8 | 1.5613 | 4 | 2.0281 |
| | 1.8213 | | |
| | 1.9352 | | |
| | 2.0828 | | |
| | 1.7574 | | |
| | 2.0421 | | |
| | 2.1864 | | |
| | 1.7970 | | |

Chapter 5

Conclusion

In this paper, the main focus was on using the RJMCMC for model selection in solving for the regression coefficients in a standard linear model, however, the selection of predictors is not known. We were given a large number of predictors and were required to select an appropriate subset to obtain the optimal model. In this problem, we focused on the first n predictors, a subset, of a total number of predictors, m .

It was found that by using RJMCMC algorithm to perform the required selection implicitly by simulating from the posterior distribution of \mathbf{b}_n , under the density, $f(\mathbf{b}_n|\mathbf{y})$. Although the computations times were relatively high in comparison to other programs/homeworks, the method was able to calculate the coefficients for the model

$$y = \sum_{i=1}^n x_i b_i + \epsilon \quad (5.1)$$